

Pliable rejection sampling

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with

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Learning the envelope for rejection sampling

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Learning the envelope for rejection sampling

Smooth functions are easier to learn

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How to adapt to the unknown smoothness?

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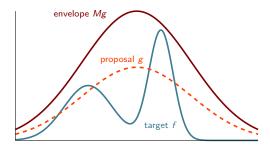
How to adapt to the unknown smoothness?

How to trade off between learning and sampling?

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Rejection Sampling

Goal: Sample from a target density f (not easy to sample from) **Tool:** Use a proposal density g (from which sampling is quite easy)



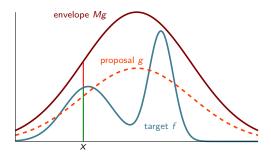
M verifies $f \leq Mg$ Rejection sampling:

- 1. Sample x from g
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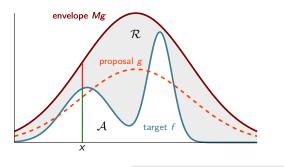
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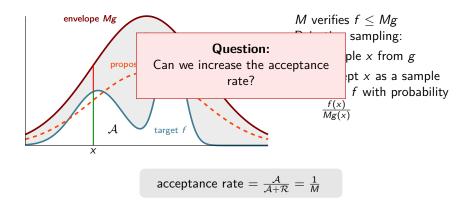
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acceptance rate =
$$\frac{A}{A+R} = \frac{1}{M}$$



Rejection Sampling

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The setting

Let $d \ge 1$ and let f be a density on \mathbb{R}^d .

Goal:

Given a number n of requests to f, what is the number T of samples Y_1, \ldots, Y_T that we can generate such that they are *i.i.d.* and sampled according to f?

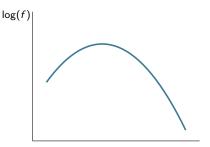
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Adaptive Rejection Sampling

Adaptive Rejection Sampling (ARS) [Gilks and Wild 1992]

- The target f is assumed to be log-concave (unimodal)
- The envelope is made of tangents at a set of points S
- ► At each rejection, the sample is added to S

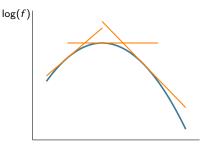




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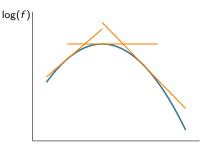




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Very strong assumption!



Improved ARS versions

Adaptive Rejection Metropolis Sampling (ARMS) [Gilks, Best and Tan 1995]

- Can deal with non-log-concave densities.
- Performs a Metropolis-Hastings control for each accepted sample
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Convex-Concave Adaptive Rejection Sampling [Gorur and Tuh 2011]

- Decomposes the target as convex + concave
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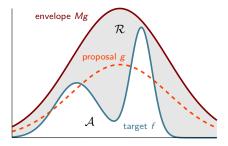
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Convexity assumption!

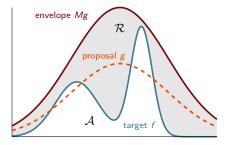


Folding the envelope



acceptance rate =
$$\frac{\mathcal{A}}{\mathcal{A} + \mathcal{R}} = \frac{1}{\widehat{M}}$$

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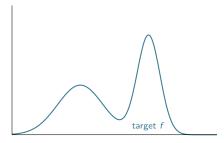


Better proposal means smaller rejection area
$${\cal R}$$

Smaller \mathcal{R} means g should have a similar "shape" to f

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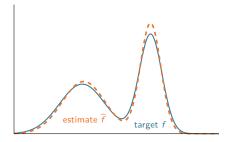


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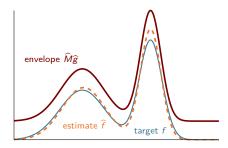
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For this purpose:

• Build an estimate \hat{f}



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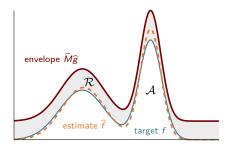
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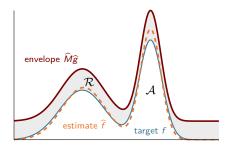
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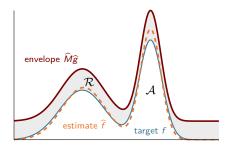
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 $1 t should be easy to sample from <math>\hat{g} \dots \hat{f} \underline{f} \underline{f}$



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- f is in a Hölder ball of smoothness s
- not very restrictive, for a small s
- ► f can be an unnormalized density (useful for some Bayesian methods)

Visualizing a 2D example

Multimodal case

$$f(x,y) \propto \left(1 + \sin\left(4\pi x - \frac{\pi}{2}\right)\right) \left(1 + \sin\left(4\pi y - \frac{\pi}{2}\right)\right)$$

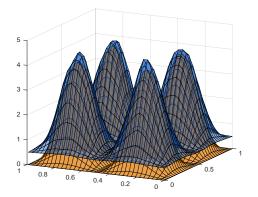


Figure: 2D target density (orange) and the pliable proposal (blue)



Pliable Rejection Sampling

Step 1: Estimating f

- f is defined on $[0, A]^d$, bounded and smooth.
- K is a positive kernel on \mathbb{R}^d (product kernel).
- Let X₁,..., X_N ∼ U_{[0,A]^d}. The (modified) kernel regression estimate is

$$\widehat{f}(x) = rac{A^d}{Nh^d} \sum_{k=1}^N f(X_i) \mathcal{K}\left(rac{X_i - x}{h}\right)$$

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Cost: N requests to f out of n.

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Assumption on the kernel K

 ${\mathcal K}_0$ be a positive univariate density kernel defined on ${\mathbb R}$

$${\cal K}=\prod_{i=1}^d {\cal K}_0$$

Furthermore, it is also of degree 2, i.e., it satisfies

$$\int_{\mathbb{R}} x K_0(x) dx = 0,$$

and, for some C' > 0

$$\int_{\mathbb{R}} x^2 K_0(x) dx \leq C'.$$

 K_0 is ε -Hölder for some $\varepsilon > 0$, i.e., $\exists C'' > 0$ s.t., for any $(x, y) \in \mathbb{R}^2$,

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Gaussian kernel satisfies this with C = 1, C' = 1, C'' = 4, and $\varepsilon = 1$

Pliable Rejection Sampling

Bounding the gap

Theorem 1

The estimate \hat{f} is such that with probability larger than $1 - \delta$, for any point $x \in [0, A]^d$,

$$\left|\widehat{f}(x) - f(x)\right| \leq H_0\left(\left(\frac{\log(NAd/\delta)}{N}\right)^{\frac{s}{2s+d}}\right)$$

where H_0 is a constant that depends on the problem parameters.

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Remaing Budget: n - N.



Pliable Rejection Sampling

Step 2: Generating Samples

• Remaining requests to f: n - N

• Let
$$r_N = A^d H_C \left(\frac{\log(NAd/\delta)}{N} \right)^{\frac{s}{2s+d}}$$

• Construct the *pliable* proposal \hat{g} out of \hat{f} :

$$\widehat{g} = \frac{\widehat{f} + r_N \, \mathcal{U}_{[0,A]^d}}{\frac{1}{N} \sum_{i=1}^N f(X_i) + r_N}$$

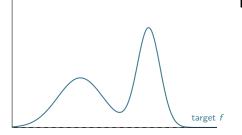
• Perform rejection sampling using \hat{g} and the empirical rejection sampling constant

$$\widehat{M} = \frac{\frac{1}{N}\sum_{i} f(X_{i}) + r_{N}}{\frac{1}{N}\sum_{i} f(X_{i}) - 5r_{N}}$$

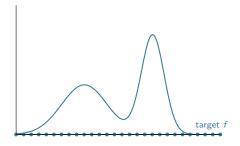


Algorithm: Pliable Rejection Sampling (PRS)

Input: *s*, *n*, δ , *H*_C



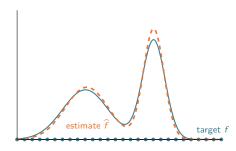




Algorithm: Pliable Rejection Sampling (PRS)

Input: *s*, *n*, δ , *H_C* **Initial Sampling** Draw uniformly at random *N* samples on $[0, A]^d$

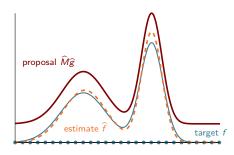




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Input: s, n, δ , H_C Initial Sampling Draw uniformly at random N samples on $[0, A]^d$ Estimation of f Estimate f using these N samples by kernel regression Generating the samples Sample n - N samples from the *pliable proposal* \hat{g} and perform Rejection

Sampling using \widehat{M} as the envelope constant

Is the sampling correct?

Theorem 1: w.p. $1 - \delta$, for any $x \in [0, A]^d$

$$\xi' \stackrel{\mathrm{def}}{=} \left| \widehat{f}(x) - f(x) \right| \leq r_N \frac{1}{A^d} = r_N \mathcal{U}_{[0,A]^d}.$$

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Hoeffding's: w.p. $1 - \delta$

$$\xi'' \stackrel{\text{def}}{=} \left\{ \left| \frac{A^d}{n} \sum_{i=1}^n f(X_i) - \int_{[0,A]^d} f(x) dx \right| \le 2A^d c \sqrt{\frac{1}{N} \log(1/\delta)} \stackrel{\text{def}}{=} c_N \right\}$$



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On, $\xi = \xi' \cap \xi''$, we have for our proposal and $8r_N \leq \int_{[0,A]^d} f(x) dx \stackrel{\text{def}}{=} m$

$$\widehat{g}^{\star} = \frac{\widehat{f} + r_{N}\mathcal{U}_{[0,A]^{d}}}{A^{d}/n\sum_{i=1}^{n}f(X_{i}) + r_{N}} \ge \frac{f}{\int_{[0,A]^{d}}f(x)dx + r_{N} + c_{N}} \\ \ge \frac{f}{\int_{[0,A]^{d}}f(x)dx}(1 - 4r_{N}/m)$$



Choice of empirical multiplication constant \widehat{M}

$$\frac{1}{1-4r_N/m} = \frac{m}{m-4r_N}$$

$$\leq \frac{A^d/N\sum_i f(X_i) + c_N}{A^d/N\sum_i f(X_i) - c_n - 4r_N}$$

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 $\widehat{M}\hat{g}^{\star}$ upperbounds f (under ξ)



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Sampling is correct whp.



$$\widehat{M} = \frac{A^d / N \sum_i f(X_i) + r_N}{A^d / N \sum_i f(X_i) - 5r_N} \le \frac{m + r_N + c_N}{m - 5r_N - c_N} \le \frac{m + 2r_N}{m - 6r_N}.$$

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On ξ , we get samples that are i.i.d. according to f, and \hat{n} will be a sum of Bernoulli random variables of parameter larger than

$$\frac{1}{\widehat{M}} \geq \frac{m - 6r_N}{m + 2r_N} \geq (1 - 6r_N/m)(1 - 4r_N/m) \geq 1 - 20r_N/m,$$

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Setting: $N = n^{\frac{2s+d}{3s+d}}$,

$$\widehat{n} \ge n \left[1 - K \log(nAd/\delta)^{\frac{s}{3s+d}} n^{-\frac{s}{3s+d}} \right].$$
(1)



A bound on the acceptance rate

The asymptotic performace

Theorem 2

Under Theorem 1's assumptions and if $H_0 < H_C$, $8r_N \leq \int_{[0,A]^d} f(x)dx$. Then, for n large enough, we have with probability larger than $1 - \delta$ that

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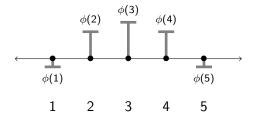
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Gumbel-Max trick: $p(i) \propto \exp(\phi(i))$ for $i \in \{1, 2, 3, 4, 5\}$

images from Chris J. Maddison

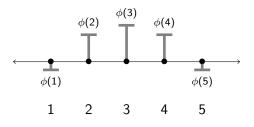




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G(1) G(2) G(3) G(4) G(5)

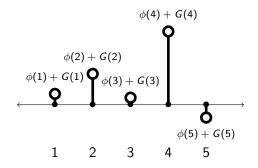


 $G(i) \sim \text{Gumbel}(0) \text{ IID}$



Gumbel-Max trick: $p(i) \propto \exp(\phi(i))$ for $i \in \{1, 2, 3, 4, 5\}$

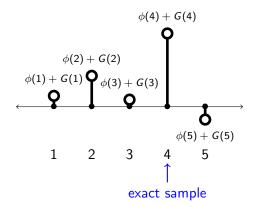
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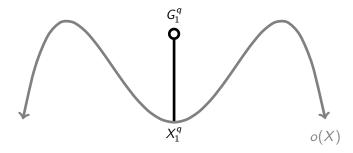
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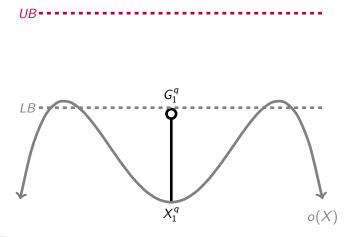


\mathbf{A}^{\star} sampling



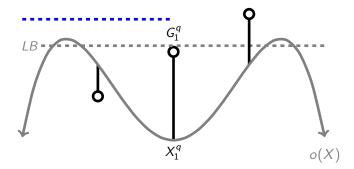


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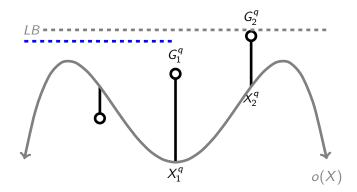


A^{*} sampling



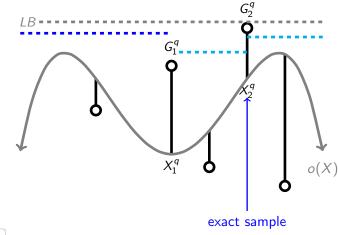
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- A^{*} needs several calls to f to generate a sample
- + PRS rejects (asymptotically) only a negligible number of samples with respect to n

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Scaling with d?



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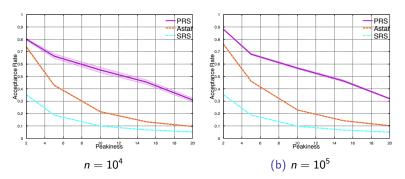
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Experiments

Scaling with peakiness



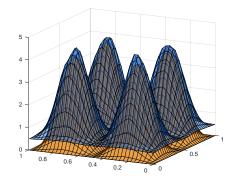
 $f \propto \frac{e^{-x}}{(1+x)^a}$, a defines the peakiness level

Figure: Acceptance rate vs. peakiness



Experiments

Two dimensional example



$n = 10^{6}$	acceptance rate	standard deviation
PRS	66.4%	0.45%
A^{\star} sampling	76.1%	0.80%
SRS	25.0%	0.01%

Table: 2D example: Acceptance rates averaged over 10 trials



Experiments

The Clutter problem

$\mathit{n}=10^{5}$, 1D	acceptance rate	standard deviation
PRS	79.5%	0.2%
A^* sampling	89.4%	0.8%
SRS	17.6%	0.1%

$n=10^5$, 2D	acceptance rate	standard deviation
PRS	51,0%	0.4%
A^{\star} sampling	56.1%	0.5%
SRS	2.10 ⁻³ %	$10^{-5}\%$

Table: Clutter problem: Acceptance rates averaged over 10 trials



Discussion

Normalized distribution

If $\int_{[0,\mathcal{A}]^d} f = 1$ then we can simplify the algorithm

$$\hat{g}^{\star} \stackrel{\text{def}}{=} \frac{1}{1+r_N} \left(\widehat{f} + r_N \mathcal{U}_{[0,A]^d} \right)$$



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Extensions for high dimensional cases (large d)

- when the mass of the distribution is localized in a few small subsets



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Thank you!

SequeL – Inria Lille GdR ISIS